Quantitative Economics for the Evaluation of the European Policy

Dipartimento di Economia e Management

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$$\overline{g_{Y/L_i}} = intercept + \beta_0 \log\left(Y/L_{i,1991}\right) + \beta_1 \bar{s}_i + \beta_2 \bar{n}_i + \beta_3 \bar{h}_i + \epsilon_i \qquad (1)$$

	Estimate	Std. Error	t-Stat.	P-value
(Intercept)	-0.0929	0.0123	-7.53	0.0000
β_0	-0.0154	0.0011	-14.57	0.0000
β_1	0.0027	0.0029	0.93	0.3532
β_2	-0.0146	0.0034	-4.31	0.0000
β_3	0.0204	0.0024	8.57	0.0000
	Res.se=0.008956 (255) DF			
	R-squared=0.6411, Adj.R-squared=0.6354			
	F-stat.=112.6 (1,255) DF, p-value=$< 2e^{-16}$			

Endogeneity in cross-region regression

Simultaneity Problem

The fact that the right-hand-side variables are not exogenous, but are **jointly determined with the growth rate** (for example the level of investment is highly correlated with growth).

- *Estimation issue*: estimates can be biased.
- *Identification issue*: the value of β can fail to illustrate how initial conditions affect expected future income differences if the saving rate is itself function of income. Hence, $\beta \geq 0$ may be compatible with at least partial convergence, while $\beta < 0$ with economic divergence if physical and human capital accumulation for rich and poor are diverging across time.

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Omitted Variables

Omitted variables appear when we would like to control for one or more additional variables but, usually because of data unavailability, we cannot include them in a regression model. $\Rightarrow$ one way to represent this situation

is to write the regression equation considering the omitted variable as part of the error term.

Instrumental Variables and Two-Stage Least Squares

Consider the linear model:

$$y = \beta_0 + \beta_1 x_1 + ... + \beta_K x_K + u \tag{2}$$

$$E(u) = 0,\ Cov(x_j, u) = 0,\ j = 1, 2, ..., K-1 \tag{3}$$

therefore x_K might be correlated with u. In other words, $x_1, ..., x_{K-1}$ are exogenous while x_K is potentially endogenous

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$\Rightarrow$ OLS estimation generally results in **inconsistent** estimators of all the β_j if $Cov(x_K, u) \neq 0$

Instrumental Variables and Two-Stage Least Squares (2)

The method of instrumental variables (IV) provides a general solution to the problem of an endogenous explanatory variable. To use the IV approach with x_K endogenous, we need an observable variable, z_1, not in equation (3) that satisfies two conditions:

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- z_1 must be uncorrelated with u: $Cov(z_1, u) = 0 \Rightarrow z_1$ is exogenous
- The second requirement involves the relationship between z_1 and the endogenous variable, x_K. Consider the regression of x_K on *all* the exogenous variables:

$$x_K = \delta_0 + \delta_1 x_1 + \delta_2 x_2 + ... + \delta_{K-1} x_{K-1} + \theta_1 z_1 + e_K \tag{4}$$

where $E(e_K) = 0$ and e_K is uncorrelated with $x_1, ..., x_{K-1}$ and z_1 $\Rightarrow \theta_1 \neq 0$

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z_1 is an **instrumental variable** candidate for x_K!

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1. Obtain the **fitted** values $\hat{x}_K$ from the regression:

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   This is called **first-stage regression**.

2. Run the OLS regression

$$y = \beta_0 + \beta_1 x_1 + ... + \beta_K \hat{x}_K + u \tag{6}$$

   This is called the **second-stage regression**, and it produces the $\hat{\beta}_j$

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- CF uses extra regressors to break the correlation between endogenous explanatory variables and unobservables affecting the dependent variable.
- The method still relies on the availability of exogenous variables that do not appear in the structural equation

Control Function (2)

Consider the linear model:

$$y = \beta_0 + \beta_1 x_1 + ... + \beta_K x_K + u \tag{7}$$

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therefore x_K might be correlated with u. In other words, $x_1, ..., x_{K-1}$ are exogenous while x_K is potentially endogenous. The *reduced form* of x_K is

the linear projection of x_K onto the exogenous variables (and the instruments):

$$x_K = \delta_0 + \delta_1 x_1 + \delta_2 x_2 + ... + \delta_{K-1} x_{K-1} + \theta_1 z_1 + e_K \tag{9}$$

with $Cov(x_j, e_K) = 0,\ j = 1, 2, ..., K-1$ and $Cov(z_1, e_K) = 0$.

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Write the linear projection of u on e_K as:

$$u = \rho e_K + \epsilon \tag{10}$$

By definition, $Cov(e_K, \epsilon) = 0$, $Cov(x_j, \epsilon) = 0$ and $Cov(z_1, \epsilon) = 0$ because u and e_K are both uncorrelated with x_j $j = 1, ..., K_1$ and z_1.

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Pluggin (9) in (6) we get:

$$y = \beta_0 + \beta_1 x_1 + ... + \beta_K x_K + \rho e_K + \epsilon \tag{11}$$

where now e_K can be viewed as an explanatory variable in the equation, and $Cov(y, \epsilon) = 0$.

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$\Rightarrow$ run OLS of y on x_j $j = 1, ..., K_1$, z_1 and e_K using random sample.

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- given that we observe $\mathbf{x}, z_1$ we can estimate the model 12 by OLS $\Rightarrow$ replace e_K with $\hat{e}_K$.

$$y = \beta_0 + \beta_1 x_1 + ... + \beta_K x_K + \rho \hat{e}_K + error \tag{13}$$

where:
$error = \epsilon + \rho(x_1, ..., x_{K-1}, z_1) \left[(\hat{\delta}_0, \hat{\delta}_1, ..., \hat{\delta}_{K-1}, \hat{\theta}) - (\delta_0, \delta_1, ..., \delta_{K-1}, \theta) \right]$
depends on the sampling error.

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$\Rightarrow$ OLS estimator of (13) will be consistent!!

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- The OLS estimate of β_j and $j = 1, ..., K$ are *identical* to the 2SLS.
- Test of endogeneity: $\rho = 0$.
- Problem: $\hat{e}_K$ is a *generated regressor* $\Rightarrow$ we need bootstrap for right standard errors!

Control Function: summaryzing

1. Obtain the **residuals** $\hat{e}_K$ from the regression:

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3. Test $\hat{\rho} = 0$.

References

- Wooldridge: Econometric Analysis of Cross Section and Panel Data; Chapter 5 and Chapter 6
- R file: endogeneity_EUregions.R in *qe4Policy_19Ott2015*.