Quantitative Economics for the Evaluation of the European Policy

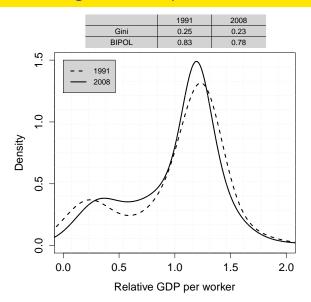
Dipartimento di Economia e Management



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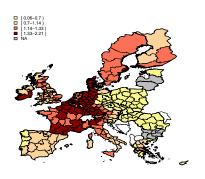
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Distribution of Regional GDP per Worker



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Geographical Distribution of Regional GDP per Worker



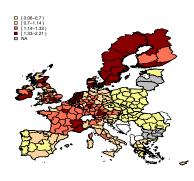
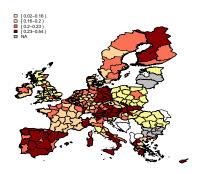


Figura: Regional GDP per Worker in 1991

Figura: Regional GDP per Worker in 2008

Geographical Distribution of Solow's Determinants



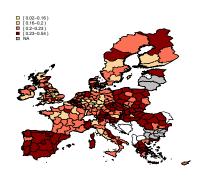
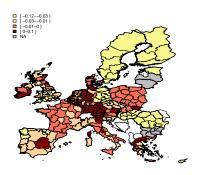


Figura: Regional Investment Rate in 1991

Figura: Regional Investment Rate in 2008

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Geographical Distribution of Solow's Determinants



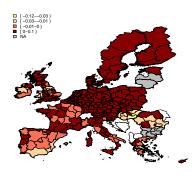
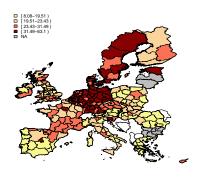


Figura: Regional Employment Growth in 1991

Figura: Regional Employment Growth in 2008

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Geographical Distribution of Solow's Determinants



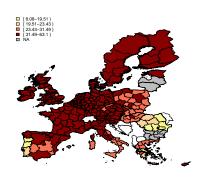


Figura: Regional Human Capital Index in 1991

Figura: Regional Human Capital Index in 2008

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Geographical Distribution of Structural and Cohesion Funds

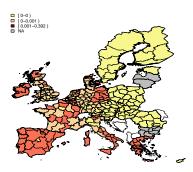


Figura: Structural and Cohesion Funds in 1989-1993

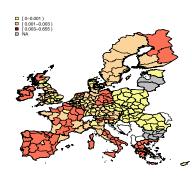


Figura: Structural and Cohesion Funds in 1994-1999

Geographical Distribution of Structural and Cohesion Funds

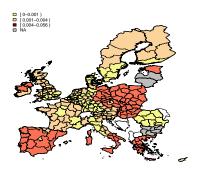


Figura: Structural and Cohesion

Funds in 2000-2006

The W Matrix

- The definition of the relation among economies in terms of their spatial structure is ambiguous.
- Differently from the time series analysis, where the notion of lagged variable is fairly unambiguous, in spatial analysis matters are complicated (Anselin, 1998).
- It must be assumed that observations are organized in spatial units, which may be points in a regular or irregular lattice, or regions on a map.

The W Matrix

Neighbours in space

Consider a variable X observed for each spatial units i = 1, ..., N. A set of **neighbours** for a spatial unit i is defined as the collection of those units j for which:

$${j|P(X_i) \neq P(X_i|X_j) \text{ and } d_{ij} \leq \epsilon_i}$$
 (1)

that is, as those locations for which the unconditional probability for X_i is different from its conditional probability given X_j .

 d_{ij} is a measure of the **distance** between i and j in a proper structured space and ϵ_i is a critical cut-off point for each spatial unit i, possibly the same.

Spatial Contiguity Matrix

The original measure of spatial autocorrelation is the **binary contiguity** between spatial units:

- the underlying structure of neighbours is expressed by 0-1 values;
- if two spatial units have a **common border** of non-zero length they are considered to be contiguous and a value 1 is assigned;
- it obviously assume the existence of a map;
- several order of contiguity may be considered.

General Spatial Weight Matrix

- To account for general measure of potential interaction between two spatial units;
- ⇒ Spatial Weigth Matrix W:
 - is based on some definition of distance as geographical distance, travel time, trade patterns etc.
 - may be symmetric or not
 - may be row-standardized

The matrix W is required because in order to address spatial autocorrelation and also model spatial interaction, we need to impose a structure to constrain the number of neighbours to be considered. This is related to Toblers first law of geography, which states that Everything depends on everything else, but closer things more so - in other words, the law implies a spatial distance decay function, such that even though all observations have an influence on all other observations, after some distance threshold that influence can be neglected.

The Moran's I

Indicators of spatial association are statistics that evaluate the existence of clusters in the spatial arrangement of a given variable.

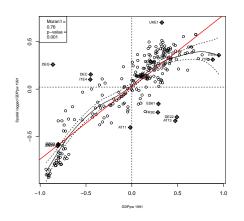
Global spatial autocorrelation is a measure of the overall clustering of the data. One of the statistics used to evaluate global spatial autocorrelation is Moran's I, defined by:

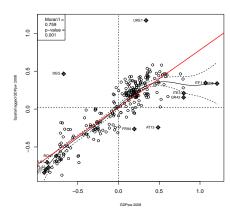
$$I = \frac{\frac{N}{S_0} \sum_i \sum_j W_{ij} Z_i Z_j}{\sum_i Z_i^2}$$
 (2)

where:

- Z_i is the deviation of the variable of interest with respect to the mean;
- W_{ij} is the matrix of weights that in some cases is equivalent to a binary matrix with ones in position i,j whenever observation i is a neighbor of observation j, and zero otherwise;
- $S_0 = \sum_i \sum_i W_{ij}$.

Negative values indicate negative spatial autocorrelation and the inverse for positive values. Values range from 1 (indicating perfect dispersion) to +1 (perfect correlation). A zero value indicates a random spatial pattern.





The LISA

Global spatial analysis or global spatial autocorrelation analysis yields only one statistic to summarize the whole study area. In other words, global analysis assumes homogeneity. If that assumption does not hold, then having only one statistic does not make sense as the statistic should differ over space.

But if there is no global autocorrelation or no clustering, we can still find clusters at a local level using local spatial autocorrelation. The fact that Moran's I is a summation of individual crossproducts is exploited by the "Local indicators of spatial association" (LISA) to evaluate the clustering in those individual units by calculating Local Moran's I for each spatial unit and evaluating the statistical significance for each I_i :

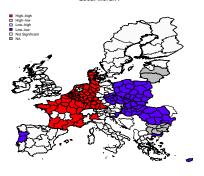
$$I_i = \frac{Z_i}{m_2} \sum_j W_{ij} Z_j \tag{3}$$

where:

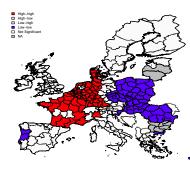
$$m_2 = \frac{\sum_i Z_i^2}{N}$$

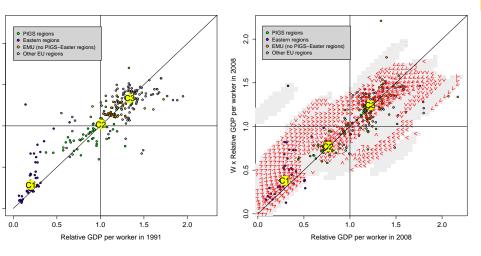
I is the Moran's I measure of global autocorrelation, I_i is local, and N is the number of analysis units in the map.

Local Moran I



Local Moran I





- lacksquare W is the row standardized spatial weight matrix
- Moran scatter plot for 1991 with the indication of three spatial clubs. The centres of three clubs of regions in 1991 are identified by k-median algorithm.
- Moran scatter plot for 2008, the three spatial clubs, and the estimated joint dynamics of (relative) GDP per worker and its spatial lagged value for the period 1991-2008 (represented by the red arrows).