Quantitative Economics for the Evaluation of the European Policy

Dipartimento di Economia e Management

Irene Brunetti ¹ Davide Fiaschi² Angela Parenti³

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¹ireneb@ec.unipi.it. ²davide.fiaschi@unipi.it. ³aparenti@ec.unipi.it. Brunetti-Fiaschi-Parenti Quant

Quantitative Economics

Distribution of Regional GDP per Worker



Quantitative Economics

Let be x a continuous random variable and f its probability density function (pdf).

The pdf characterizes the distribution of the random variable x since it tells "how x is distributed".

Moreover, from pdf it is possible to calculate the mean and the variance (it they exists) of x and the probability that x takes on values in a given interval.

Histograms are nonparametric estimates of an *unknown density function*, f(x), **without assuming any well-known functional form**. In order to build an histogram, you have to:

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2 count how many observations fall into each bin $(n_i$ for each bin j;

• for each bin divide the frequency by the sample size *n* and the binwidth *h*, to get the relative frequencies $f_j = \frac{n_j}{nh}$

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Histogram: Cont.



Brunetti-Fiaschi-Parenti

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 \rightarrow we need to find an "optimal" binwidth, which represents an optimal compromise.

Problems with the histogram:

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Problems with the histogram:

• each observation x in $[m_j - \frac{h}{2}, m_j + \frac{h}{2}]$ is estimated by the same value, $\hat{f}_h(m_j)$, where m_j is the center of the bin;

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- f(x) is estimated using the observations that fall in the interval containing x, and that receive the same weight in the estimation. That is, for x ∈ B_i,

$$\hat{f}_h(m_j) = \frac{1}{nh} \sum_{i=1}^n I(X_i \in B_j),$$

where I is the indicator function.

Nonparametric density estimation

• Density estimation is a generalization of the histogram.

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Nonparametric density estimation

- Density estimation is a generalization of the histogram.
- It is based on **Kernel functions**: estimate f(x) using the observations that fall into an interval around x, which (typically) receive decreasing weight the further they are from x.

Consider the *uniform* kernel function, which assigns the same weight to all observations in an interval of length 2h around observation x, [x - h, x + h):

$$\hat{f}_h(x) = \frac{1}{2nh} \sharp \{ X_i \in [x-h, x+h) \}$$

can be obtained by means of a kernel function K(u) such that:

$$K(u)=\frac{1}{2}I(|u|\leq 1)$$

where I is the indicator function and $u = (x - X_i)/h$.

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- Each observation that falls into the interval [x h, x + h] the indicator function takes on value 1
- Each contribution to the function is weighted equally no matter how close the observation X_i is to x

Kernel functions: Cont.

A Kernel function in general (e.g. Epanechnikov, Gaussian, etc), assigns higher weights to observations in [x - h, x + h) closer to x.



Kernel density

A kernel density estimation appears as a sum of bumps: at a given x, the value of $\hat{f}_h(x)$ is found by vertically summing over the "bumps":



$$\hat{f}_h(x) = \sum_{i=1}^n \frac{1}{nh} \mathcal{K}\left(\frac{x - X_i}{h}\right) = \sum_{i=1}^n \frac{1}{n} \underbrace{\mathcal{K}_h(x - X_i)}_{\text{"rescaled kernel function"}} \underbrace{\mathcal{K}_h(x - X_i)}_{\text{"rescaled kernel function"}}$$

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Quantitative Economics

Same problems found for the histogram, that is the bias and the variance depending on h:

$$Bias\{\hat{f}_{h}(x)\} = E\{\hat{f}_{h}(x)\} - f(x);$$

that positively depends on h;

$$Var{\hat{f}_h(x)} = Var\left\{\sum_{i=1}^n \frac{1}{n}K_h(x-X_i)\right\};$$

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So, how do we choose h given the trade-off between bias and variance?

(a) Define MSE (mean squared error)

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- (c) Define AMISE (an approximation of MISE) \rightarrow still h_{opt} depends on the unknown f(x), in particular on its second derivative f''(x).
- (d) One possibility is a plug-in method suggested by Silverman, and consists in assuming that the unknown function is a Gaussian density function (whose variance is estimated by the sample variance). In this case h_{opt} has a simple formulation, and can be defined as a rule-of-thumb bandwidth.

Brunetti-Fiaschi-Parenti

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- get an initial estimate to have a rough idea of the density
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- Step 2 Define local bandwidth factor λ_i by:

$$\lambda_i = [f(\tilde{x_i})/g]^{-\alpha} \tag{1}$$

where g is the geometric mean of the $\tilde{f}(x_i)/$, $logg = n^{-1} \sum log \tilde{f}(x_i)$; and α the sensitivity parameter ($\alpha \leq 0$)

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where g is the geometric mean of the $\tilde{f}(x_i)/$, $logg = n^{-1} \sum log \tilde{f}(x_i)$; and α the sensitivity parameter ($\alpha \leq 0$)

Step 3 Define the adaptive kernel estimate $\hat{f}(x)$ by:

$$\hat{f}(x) = nh^{-1} \sum \lambda_i^{-1} \mathcal{K}\{h^{-1}\lambda_i^{-1}(x-X_i)\}$$
(2)

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- Select B independent bootstrap samples {X*1, ..., X*B}, each consisting of n data values drawn with replacement from x.

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- Estimate from sample x the density \hat{f} .
- Select *B* independent bootstrap samples $\{X^{*1}, ..., X^{*B}\}$, each consisting of *n* data values drawn with replacement from *x*.
- Solution Estimate the density \hat{f}_b^* corresponding to each bootstrap sample b = 1, ..., B.

Given a sample of observations $X = \{X_1, ..., X_m\}$ where each X_i is a vector of dimension *n* the bootstrap algorithm is the following.

- Estimate from sample x the density \hat{f} .
- Select B independent bootstrap samples {X*1,...,X*B}, each consisting of n data values drawn with replacement from x.
- Solution Estimate the density \hat{f}_b^* corresponding to each bootstrap sample b = 1, ..., B.

The distribution of \hat{f}^* about \hat{f} can therefore be used to mimic the distribution of \hat{f} about f, that is it can be used to calculate the confidence intervals for estimates.

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- Estimate: Chapter 1
- Inference (confidence bands): Chapter 2

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• Estimate: Chapter 5.3

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